

Exam of 26.01.2000 — Solutions

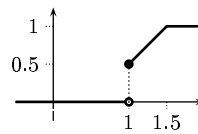
1

$$X = \begin{cases} 1 & \text{if } T_1 > 1, \\ 1 + T_2 & \text{if } T_1 < 1. \end{cases}$$

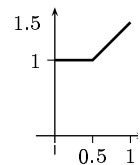
1a

$$\begin{aligned} \mathbb{P}(X = 1) &= \mathbb{P}(T_1 > 1) = \frac{1}{2}; \\ \mathbb{P}(1 < X < x) &= \mathbb{P}(1 + T_2 < x, T_1 < 1) = \mathbb{P}(T_2 < x - 1) \mathbb{P}(T_1 < 1) = \\ &= 2(x - 1) \cdot \frac{1}{2} = x - 1 \quad \text{for } x \in (1, 1.5); \end{aligned}$$

$$F_X(x) = \begin{cases} 0 & \text{for } x < 1, \\ x - 0.5 & \text{for } 1 \leq x \leq 1.5; \\ 1 & \text{for } x \geq 1.5. \end{cases}$$



$$X^*(p) = \begin{cases} 1 & \text{for } 0 < p < 0.5; \\ 0.5 + p & \text{for } 0.5 < p < 1. \end{cases}$$



1b

$$x_{1/2} = X^*\left(\frac{1}{2}\right) = 1; \quad x_{1/4} = X^*\left(\frac{1}{4}\right) = 1; \quad x_{3/4} = X^*\left(\frac{3}{4}\right) = 1.25.$$

1c

The support: $[1, 1.5]$. Discrete, singular, and absolutely continuous components:

$$\begin{aligned} p_d &= \frac{1}{2}, \quad p_s = 0, \quad p_{ac} = \frac{1}{2}; \\ F_{X,d}(x) &= \begin{cases} 0 & \text{for } x < 1, \\ 1 & \text{for } x \geq 1; \end{cases} \\ F_{X,ac}(x) &= \begin{cases} 0 & \text{for } x \leq 1, \\ 2(x - 1) & \text{for } 1 \leq x \leq 1.5, \\ 1 & \text{for } x \geq 1.5; \end{cases} \\ f_{X,ac}(x) &= 2 \cdot \mathbf{1}_{(1,1.5)}(x). \end{aligned}$$

1d

$$\begin{aligned} \mathbb{E}X &= \int_0^1 X^*(p) dp = \int_0^{0.5} dp + \int_{0.5}^1 (0.5 + p) dp = 0.5 + 0.625 = 1.125; \\ \mathbb{E}X &= p_d \mathbb{E}(X | X \in A_X) + p_{ac} \mathbb{E}(X | X \notin A_X) = \frac{1}{2} \cdot 1 + \frac{1}{2} \cdot 1.25 = 1.125. \end{aligned}$$

(Other ways may be used.)

1e

$$\mathbb{E}(1.5 - X)^{-\alpha} = p_d \int (1.5 - x)^{-\alpha} dF_{X,d}(x) + p_{ac} \int (1.5 - x)^{-\alpha} f_{X,ac}(x) dx;$$

the discrete part converges for all α (since the atom is not at 1.5); only $(1.5 - x)^{-\alpha} f_{X,ac}(x)$ must be integrable when $x = 1.5 - \varepsilon, \varepsilon \rightarrow 0+$. So, $\varepsilon^{-\alpha}$ must be integrable, which holds if and only if $\alpha < 1$.

3 _____

We have two cases: $T_1 > 1$ and $T_1 < 1$.

3a

Case $T_1 > 1$:

$X = 1, Y = T_1 + T_2, T_1 \in (1, 2), T_2 \in (0, 0.5)$; a singular 2-dim distribution, with the support $\{1\} \times [1, 2.5]$ (a segment on the plane).

Case $T_1 < 1$:

$X = 1 + T_2, Y = T_1 + T_2, T_1 \in (0, 1), T_2 \in (0, 0.5)$; an absolutely continuous 2-dim distribution with the support $\{(x, y) : x \in [1, 1.5], y \in [x - 1, x]\}$ (a parallelogram).

So, $p_d = 0, p_s = 1/2, p_{ac} = 1/2$. The support is the union of the segment and the parallelogram.

3b

$$\begin{aligned} \mathbb{E}(XY | T_1 > 1) &= \mathbb{E}(Y | T_1 > 1) = \mathbb{E}(T_1 + T_2 | T_1 > 1) = \\ &= \mathbb{E}(T_1 | T_1 > 1) + \mathbb{E}(T_2 | T_1 > 1) = 1.5 + 0.25 = 1.75; \\ \mathbb{E}(XY | T_1 < 1) &= \mathbb{E}((1 + T_2)(T_1 + T_2) | T_1 < 1) = \\ &= \mathbb{E}(T_1 | T_1 < 1) + \mathbb{E}(T_2 | T_1 < 1) + \mathbb{E}(T_1 T_2 | T_1 < 1) + \mathbb{E}(T_2^2 | T_1 < 1) = \\ &= 0.5 + 0.25 + 0.5 \cdot 0.25 + \frac{1}{12} \approx 0.958; \quad (\text{or } \frac{23}{24}) \\ \mathbb{E}(XY) &= \mathbb{E}(XY | T_1 > 1) \mathbb{P}(T_1 > 1) + \mathbb{E}(XY | T_1 < 1) \mathbb{P}(T_1 < 1) \approx \\ &\approx 1.75 \cdot 0.5 + 0.958 \cdot 0.5 \approx 1.354; \\ \text{Cov}(X, Y) &\approx 1.354 - 1.125 \cdot 1.25 \approx -0.052. \quad (\text{or } -\frac{5}{96}) \end{aligned}$$

So, X and Y are correlated, therefore, dependent.

3c

$P_{X,Y,ac}$ is $P_{X,Y|T_1 < 1}$, the uniform distribution on the parallelogram. The area of the parallelogram is $1/2$, so, the density is

$$f_{X,Y,ac}(x, y) = \begin{cases} 2 & \text{if } x \in [1, 1.5], y \in [x - 1, x], \\ 0 & \text{otherwise.} \end{cases}$$

4 _____

4a

Yes, $a_n \leq a_{n+1}$, since $\max(X_1, \dots, X_n) \leq \max(X_1, \dots, X_{n+1})$. The equality $a_n = a_{n+1}$ is possible, since the distribution may be degenerate, $X_1 = \text{const}$.

4b

If X_1 is bounded from above, that is, $\mathbb{P}(X_1 \leq c) = 1$ for some c , then the sequence (a_n) is bounded (by c). Conversely, if (a_n) is bounded, then $\lim_n \mathbb{E} \max(X_1, \dots, X_n) < \infty$, which implies (by monotone convergence theorem) that $\lim_n \max(X_1, \dots, X_n) < \infty$ almost sure, that is, $\sup(X_1, X_2, \dots) < \infty$ almost sure. That is possible only for bounded X_1 . Otherwise, for every x the inequality $X_n > x$ holds infinitely often.

4c

$$\frac{a_n}{n} = \frac{1}{n} \mathbb{E} \max(X_1, \dots, X_n) = \frac{1}{n} \left(- \int_{-\infty}^0 F^n(x) dx + \int_0^{\infty} (1 - F^n(x)) dx \right).$$

The first term tends to 0; consider the second term. We have

$$1 - F^n(x) = (1 - F(x))(1 + F(x) + F^2(x) + \dots + F^{n-1}(x)) \leq n(1 - F(x)).$$

Therefore for every M ,

$$\begin{aligned} \frac{1}{n} \int_0^{\infty} (1 - F^n(x)) dx &= \underbrace{\frac{1}{n} \int_0^M (1 - F^n(x)) dx}_{\rightarrow 0} + \frac{1}{n} \int_M^{\infty} (1 - F^n(x)) dx; \\ \frac{1}{n} \int_M^{\infty} (1 - F^n(x)) dx &\leq \int_M^{\infty} (1 - F(x)) dx \xrightarrow{M \rightarrow \infty} 0, \end{aligned}$$

since

$$\int_M^{\infty} (1 - F(x)) dx = \int_0^{\infty} (1 - F(x)) dx - \underbrace{\int_0^M (1 - F(x)) dx}_{\rightarrow \int_0^{\infty} (1 - F(x)) dx}$$

A more elegant solution found by a student on the exam:

Denote $Y_n = \max(X_1, \dots, X_n)$, then $F_{Y_n}(\cdot) = F_{X_1}^n(\cdot)$, thus $Y_n^*(p) = X_1^*(\sqrt[n]{p})$ and $a_n = \mathbb{E}Y_n = \int_0^1 Y_n^*(p) dp = \int_0^1 X_1^*(\sqrt[n]{p}) dp = \int_0^1 X_1^*(u) \cdot nu^{n-1} du$; so, $\frac{1}{n}a_n = \int_0^1 X_1^*(u)u^{n-1} du \rightarrow 0$ by the monotone convergence theorem.

5

5a

$$F_{X|V=1}(x) = F_{U|V=1}(x) = F_U(x) = \int_0^x f_U(u) du;$$

$$F_{X|V=2}(x) = F_{2U|V=2}(x) = F_{2U}(x) = F_U\left(\frac{x}{2}\right) = \int_0^{x/2} f_U(u) du = \frac{1}{2} \int_0^x f_U\left(\frac{u}{2}\right) du;$$

$$f_{X|V=1}(x) = f_U(x);$$

$$f_{X|V=2}(x) = \frac{1}{2}f_U\left(\frac{x}{2}\right).$$

5b

$$\mathbb{P}(V = 1 | X = x) = \frac{f_{X|V=1}(x)\mathbb{P}(V = 1)}{f_X(x)} = \frac{f_U(x) \cdot \frac{1}{2}}{f_U(x) \cdot \frac{1}{2} + \frac{1}{2}f_U\left(\frac{x}{2}\right) \cdot \frac{1}{2}} = \frac{2f_U(x)}{2f_U(x) + f_U\left(\frac{x}{2}\right)};$$

$$\mathbb{P}(V = 2 | X = x) = \frac{f_U\left(\frac{x}{2}\right)}{2f_U(x) + f_U\left(\frac{x}{2}\right)}.$$

5c

$$\mathbb{E}(Y | X = x) = \mathbb{E}\left(\frac{3-V}{V}X | X = x\right) = x \mathbb{E}\left(\frac{3-V}{V} | X = x\right) =$$

$$= x\left(2\mathbb{P}(V = 1 | X = x) + \frac{1}{2}\mathbb{P}(V = 2 | X = x)\right) =$$

$$= x\left(2\frac{2f_U(x)}{2f_U(x) + f_U\left(\frac{x}{2}\right)} + \frac{1}{2}\left(1 - \frac{2f_U(x)}{2f_U(x) + f_U\left(\frac{x}{2}\right)}\right)\right) = x\left(\frac{3}{2}\frac{2f_U(x)}{2f_U(x) + f_U\left(\frac{x}{2}\right)} + \frac{1}{2}\right).$$

5d

$\mathbb{E}(Y | X) > X$ means $\frac{3}{2}\frac{2f_U(x)}{2f_U(x)+f_U(\frac{x}{2})} + \frac{1}{2} > 1$, that is, $6f_U(x) > 2f_U(x) + f_U(\frac{x}{2})$, that is, $4f_U(x) > f_U(\frac{x}{2})$. That must hold for almost all x .

Yes, such f_U exist. For example, $f_U(x) = \text{const} \cdot \min(1, 1/x^\alpha)$ for $\alpha \in (1, 2)$.

If X, Y are integrable, then $\mathbb{E}(Y | X) > X$ implies $\mathbb{E}Y > \mathbb{E}X$, since $\mathbb{E}(\mathbb{E}(Y | X)) = \mathbb{E}Y$. However, we have $\mathbb{E}(X) = \mathbb{E}(Y)$ (these are identically distributed), so, $\mathbb{E}(Y | X) > X$ implies that $\mathbb{E}(X) = \infty = \mathbb{E}Y$.