

## 2 Markov and strong Markov

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### 2a Restart at a nonrandom time

Let  $X$  be any one of the three processes introduced in Sect. 1 (the Brownian motion, the Cauchy process, the special Lévy process) on a probability space  $(\Omega, \mathcal{F}, P)$ . We construct a random function  $Y$  on the product  $\Omega^2 = \Omega \times \Omega$  (that is,  $(\Omega, \mathcal{F}, P) \times (\Omega, \mathcal{F}, P)$ ) by glueing together two independent sample functions as follows:

$$(2a1) \quad Y(t)(\omega_1, \omega_2) = \begin{cases} X(t)(\omega_1) & \text{for } t \leq 1, \\ X(1)(\omega_1) + X(t-1)(\omega_2) & \text{for } t \geq 1. \end{cases}$$

Clearly, sample functions of  $Y$  are right continuous.

**2a2 Exercise.**  $Y$  is distributed like  $X$ .<sup>1</sup>

Prove it.

This is the *Markov property*: at the instant 1 the process  $X$  forgets its past and retains only a single point,  $X(1)$ .<sup>2</sup> Of course, the Markov property holds at every instant  $t \in (0, \infty)$ , not just 1.

We turn to the Brownian motion,  $B$ . Given  $x \in (0, \infty)$ , we define the *hitting time*  $T_x : \Omega \rightarrow [0, \infty]$  by

$$(2a3) \quad T_x = \inf\{t : B(t) = x\}$$

(as usual,  $\inf \emptyset = \infty$ ).

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<sup>1</sup>Recall 1c, especially 1c3. See also 2f4.

<sup>2</sup>By the way, a process with differentiable sample functions cannot be Markov (unless it is nonrandom); it have to retain  $X'(1)$ .

**2a4 Exercise.** (a)  $T_x$  is measurable (in  $\omega$ , for a fixed  $x$ ); (b) the distribution of  $T_x$  is uniquely determined, that is, does not depend on the choice of  $(\Omega, \mathcal{F}, P)$  and  $B$  as far as  $B$  is a Brownian motion.<sup>1</sup>

Prove it.

Such statements should be made every time we construct a random variable out of the Brownian motion;<sup>2</sup> however, they will be usually omitted.

**2a5 Exercise.**  $T_x$  is distributed like  $x^2 T_1$ .

Prove it.

We introduce the random variable<sup>3</sup>

$$(2a6) \quad L = \max\{t \in [0, 1] : B(t) = 0\},$$

and want to calculate its distribution,

$$\mathbb{P}(L < t) = \mathbb{P}(\forall s \in [t, 1] \ B(s) \neq 0) = ?$$

Given  $B(t) = x > 0$ , the conditional probability of this event should be equal to

$$\mathbb{P}(\forall s \in [0, 1-t] \ B(s) \neq x) = \mathbb{P}(T_x > 1-t) = \mathbb{P}\left(T_1 > \frac{1-t}{x^2}\right)$$

(think, why); for  $x < 0$  the situation is similar. We guess that

$$(2a7) \quad \mathbb{P}(L < t) = \int_{-\infty}^{\infty} p_t(x) \mathbb{P}\left(T_1 > \frac{1-t}{x^2}\right) dx,$$

where  $p_t(x) = (2\pi t)^{-1/2} \exp(-\frac{x^2}{2t})$ .

The proof combines the Markov property of the Brownian motion with the Fubini theorem. We use  $\omega_1$  on  $[0, t]$ , switch to  $\omega_2$  on  $[t, 1]$ , substitute this combination for  $B$  into  $L$  and get

$$\begin{aligned} \mathbb{P}(L < t) &= \mathbb{P}(\forall s \in [t, 1] \ B(s) \neq 0) = \\ &= (P \times P)\{(\omega_1, \omega_2) : \forall s \in [t, 1] \ B(t)(\omega_1) + B(s-t)(\omega_2) \neq 0\} = \\ &= \int_{\Omega} f(B(t)(\omega_1)) P(d\omega_1) = \mathbb{E} f(B(t)) = \int_{\mathbb{R}} p_t(x) f(x) dx, \end{aligned}$$

<sup>1</sup>See also 2f3, 2f4.

<sup>2</sup>For instance,  $L$  and  $R$ , see (2a6), (2a9).

<sup>3</sup>' $L$  is for left or last' [1, Sect. 7.2, Exer. 2.2].

where

$$\begin{aligned} f(x) &= P\{\omega_2 : \forall s \in [t, 1] \ x + B(s-t)(\omega_2) \neq 0\} = \\ &= \mathbb{P}(\forall s \in [0, 1-t] \ B(s) \neq -x) = \mathbb{P}(T_{|x|} > 1-t) = \mathbb{P}\left(T_1 > \frac{1-t}{x^2}\right); \end{aligned}$$

(2a7) follows.

**2a8 Exercise.** Let<sup>1</sup>

$$(2a9) \quad R = \inf\{t \in [1, \infty) : B(t) = 0\}$$

(possibly,  $\infty$ ).<sup>2</sup> Then

$$(2a10) \quad \mathbb{P}(R > 1+t) = \int_{-\infty}^{\infty} p_1(x) \mathbb{P}\left(T_1 > \frac{t}{x^2}\right) dx.$$

Prove it.

## 2b Hit and restart

Similarly to (2a1) we let (recall (2a3))

$$(2b1) \quad Y(t)(\omega_1, \omega_2) = \begin{cases} B(t)(\omega_1) & \text{for } t \leq T_1(\omega_1), \\ 1 + B(t - T_1(\omega_1))(\omega_2) & \text{for } t \geq T_1(\omega_1). \end{cases}$$

**2b2 Proposition.**  $Y$  is distributed like  $B$ .

The proof will be given in 2c, but do not hesitate to use 2b2 now.

This is a special case of *strong Markov property*.<sup>3</sup>

You see, the process  $B$  forgets the past when hitting the level 1. Of course, the same happens when hitting  $x$ , for every  $x \in \mathbb{R}$ , not just 1.

**2b3 Exercise.** Prove that

$$\mathbb{P}\left(\max_{[0,t]} B(\cdot) \geq 1\right) = 2\mathbb{P}(B(t) \geq 1).$$

Similarly,  $\mathbb{P}(\max_{[0,t]} B(\cdot) \geq x) = 2\mathbb{P}(B(t) \geq x)$  for all  $x \in [0, \infty)$ . Thus,

$$(2b4) \quad \max_{[0,t]} B(\cdot) \text{ is distributed like } |B(t)|.$$

<sup>1</sup>' $R$  is for right or return' [1, Sect. 7.2, Exer. 2.1].

<sup>2</sup>But see 2b5.

<sup>3</sup>See also 2f8.

The distribution of  $T_x$  is therefore

$$\begin{aligned}\mathbb{P}(T_x \leq t) &= \mathbb{P}\left(\max_{[0,t]} B(\cdot) \geq x\right) = 2\mathbb{P}(B(t) \geq x) = \\ &= 2\mathbb{P}\left(B(1) \geq \frac{x}{\sqrt{t}}\right) = 2 \int_{x/\sqrt{t}}^{\infty} p_1(y) \, dy.\end{aligned}$$

**2b5 Exercise.** Prove that

$$\inf_{[0,\infty)} B(\cdot) = -\infty, \quad \sup_{[0,\infty)} B(\cdot) = \infty \quad \text{a.s.}$$

**2b6 Exercise.** Almost surely,

$$\forall \varepsilon > 0 \quad \left( \min_{[0,\varepsilon]} B(\cdot) < 0 \text{ and } \max_{[0,\varepsilon]} B(\cdot) > 0 \right).$$

Prove it.

**2b7 Exercise.**  $B$  does not restart at the random time  $L$  (defined by (2a6)).

Prove it.

Now we are in position to finalize the calculation of the distribution of  $L$  and  $R$  started in (2a7), (2a10); the integrals need some effort, and give

$$(2b8) \quad \mathbb{P}(L \leq t) = \frac{2}{\pi} \arcsin \sqrt{t} \quad \text{for } 0 \leq t \leq 1,$$

$$(2b9) \quad \mathbb{P}(R \leq t) = \frac{2}{\pi} \arctan \sqrt{t-1} \quad \text{for } 1 \leq t < \infty,$$

see [1, Sect. 7.4, Example 4.4].

Let us calculate the density:

$$\begin{aligned}\frac{d}{dt} \mathbb{P}(T_x \leq t) &= 2 \frac{d}{dt} \int_{x/\sqrt{t}}^{\infty} p_1(y) \, dy = \\ &= -2p_1\left(\frac{x}{\sqrt{t}}\right) \cdot x \cdot \left(-\frac{1}{2}t^{-3/2}\right) = \frac{x}{t^{3/2}} p_1\left(\frac{x}{\sqrt{t}}\right) = \\ &= \frac{x}{t} p_t(x) = \frac{1}{\sqrt{2\pi}} \frac{x}{t^{3/2}} \exp\left(-\frac{x^2}{2t}\right);\end{aligned}$$

the derivative is continuous on  $[0,\infty)$  (in spite of  $t$  in the denominator); we got the density (of the distribution) of  $T_x$ . Note that  $\mathbb{E}T_x = \infty$ . Interestingly,  $T_x$  is distributed like the special Lévy process at *time*  $x$ .

**2b10 Exercise.** For all  $x, y \in (0, \infty)$ ,

$T_{x+y} - T_x$  is independent of  $T_x$  and distributed like  $T_y$ .

Prove it.

The formula  $p_{s+t} = p_s * p_t$  for  $p_t(x) = \frac{t}{\sqrt{2\pi x^3/2}} \exp(-\frac{t^2}{2x})$ , claimed in 1a without proof, follows from 2b10!

Similarly to 2b10, the process  $(T_x)_{x \in [0, \infty)}$  has stationary independent increments. Also, its sample functions are continuous from the left (think, why).

The random function  $(T_x)_{x \in [0, \infty)}$  is distributed as the left-continuous modification of the special Lévy process.

See also [1, Sect. 7.4].

But wait, we did not prove 2b2 yet...<sup>1</sup>

## 2c Delayed restart

An important step toward the proof of Prop. 2b2 is made here. Instead of the random time  $T_1$  taking on a continuum of values we introduce (for a given  $n$ ) a random time  $\tau_n$  with a finite number of values,

$$(2c1) \quad \tau_n = \frac{k}{2^n} \quad \text{whenever} \quad \frac{k-1}{2^n} < T_1 \leq \frac{k}{2^n} \quad \text{for } k = 1, 2, \dots, 2^{2^n};$$

$$\tau_n = \infty \quad \text{whenever } T_1 > 2^n.$$

Clearly,  $\tau_n \downarrow T_1$  a.s., as  $n \rightarrow \infty$ .

Similarly to (2b1) we restart at  $\tau_n$ ,

$$(2c2) \quad Y_n(t)(\omega_1, \omega_2) = \begin{cases} B(t)(\omega_1) & \text{for } t \leq \tau_n(\omega_1), \\ B(\tau_n(\omega_1))(\omega_1) + B(t - \tau_n(\omega_1))(\omega_2) & \text{for } t \geq \tau_n(\omega_1). \end{cases}$$

Similarly to 2b2 we claim the following.

**2c3 Lemma.** For every  $n$  the random function  $Y_n$  is distributed like  $B$ .

The proof will be given in 2e. Now we'll deduce 2b2 from 2c3.

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<sup>1</sup>“It may be difficult for the novice to appreciate the fact that twenty five years ago a formal proof of the strong Markov property was a major event.” Kai Lai Chung, John B. Walsh, “Markov processes, Brownian motion, and time symmetry”, second edition, Springer (1982 and) 2005; see page 73.

*Proof of 2b2 (assuming 2c3).* The random function  $Y$  defined by (2b1) is evidently continuous. In order to prove that  $Y$  is distributed like  $B$  it is sufficient to check that  $(Y(t_1), \dots, Y(t_j)) \sim (B(t_1), \dots, B(t_j))$  for all  $j$  and  $t_1, \dots, t_j \in (0, \infty)$ . To this end it is sufficient to check that

$$(2c4) \quad \mathbb{E} \varphi(Y(t_1), \dots, Y(t_j)) = \mathbb{E} \varphi(B(t_1), \dots, B(t_j))$$

for every  $j$  and every bounded continuous  $\varphi : \mathbb{R}^j \rightarrow \mathbb{R}$ .

By 2c3,

$$(2c5) \quad \mathbb{E} \varphi(Y_n(t_1), \dots, Y_n(t_j)) = \mathbb{E} \varphi(B(t_1), \dots, B(t_j))$$

for all  $n$ . As  $n \rightarrow \infty$ , we have (for almost all  $\omega_1, \omega_2$ )

$$\begin{aligned} \tau_n(\omega_1) &\downarrow T_1(\omega_1); \\ B(\tau_n(\omega_1))(\omega_1) &\rightarrow B(T_1(\omega_1))(\omega_1) = 1; \\ t - \tau_n(\omega_1) &\rightarrow t - T_1(\omega_1); \\ B(t - \tau_n(\omega_1))(\omega_2) &\rightarrow B(t - T_1(\omega_1))(\omega_2); \\ Y_n(t)(\omega_1, \omega_2) &\rightarrow Y(t)(\omega_1, \omega_2) \end{aligned}$$

for  $t \geq T_1(\omega_1)$ . And clearly  $Y_n(t)(\omega_1, \omega_2) = B(t)(\omega_1) = Y(t)(\omega_1, \omega_2)$  for  $t < T_1(\omega_1)$ , if  $n$  is large enough. Thus,  $Y_n(t) \rightarrow Y(t)$  a.s. (for each  $t$ ); therefore

$$\mathbb{E} \varphi(Y_n(t_1), \dots, Y_n(t_j)) \rightarrow \mathbb{E} \varphi(Y(t_1), \dots, Y(t_j))$$

by the bounded convergence theorem. In combination with (2c5) it gives (2c4).  $\square$

## 2d Maybe restart, maybe not

Here we prove Lemma 2c3 for the simplest case,  $n = 0$ . (Be careful, mind 2b7!) By (2c1),

$$\tau_0 = \begin{cases} 1 & \text{if } T_1 \leq 1, \\ \infty & \text{if } T_1 > 1. \end{cases}$$

By (2c2),<sup>1</sup>

$$Y_0(t)(\omega_1, \omega_3) = \begin{cases} B(t)(\omega_1) & \text{if } t \leq 1, \\ B(t)(\omega_1) & \text{if } t \geq 1 \text{ and } \max_{[0,1]} B(\cdot)(\omega_1) < 1, \\ B(1)(\omega_1) + B(t-1)(\omega_3) & \text{if } t \geq 1 \text{ and } \max_{[0,1]} B(\cdot)(\omega_1) \geq 1. \end{cases}$$

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<sup>1</sup>Why  $\omega_3$ ? Wait a little...

We want to prove that  $Y_0 \sim B$ . The distribution of  $Y_0$  does not change if we replace  $B$  with another process  $X$  distributed like  $B$ . We choose (recall (2a1))

$$X(t)(\omega_1, \omega_2) = \begin{cases} B(t)(\omega_1) & \text{for } t \leq 1, \\ B(1)(\omega_1) + B(t-1)(\omega_2) & \text{for } t \geq 1 \end{cases}$$

and consider

$$Y(t)(\omega_1, \omega_2, \omega_3) = \begin{cases} X(t)(\omega_1, \omega_2) & \text{if } t \leq 1, \\ X(t)(\omega_1, \omega_2) & \text{if } t \geq 1 \text{ and } \max_{[0,1]} X(\cdot)(\omega_1, \omega_2) < 1, \\ X(1)(\omega_1, \omega_2) + B(t-1)(\omega_3) & \text{if } t \geq 1 \text{ and } \max_{[0,1]} X(\cdot)(\omega_1, \omega_2) \geq 1. \end{cases}$$

Similarly to 2a4,<sup>1</sup>  $Y$  is distributed like  $Y_0$ . We have

$$Y(t)(\omega_1, \omega_2, \omega_3) = \begin{cases} B(t)(\omega_1) & \text{if } t \leq 1, \\ B(1)(\omega_1) + B(t-1)(\omega_2) & \text{if } t \geq 1 \text{ and } \omega_1 \in A, \\ B(1)(\omega_1) + B(t-1)(\omega_3) & \text{if } t \geq 1 \text{ and } \omega_1 \notin A, \end{cases}$$

where  $A = \{\omega_1 : \max_{[0,1]} B(\cdot)(\omega_1) < 1\}$ .

**2d1 Exercise.** Let  $(\Omega, \mathcal{F}, P)$  be a probability space,  $A \subset \Omega$  a measurable set,  $f : \Omega^2 \rightarrow \mathbb{R}$  a bounded measurable function. Define  $g : \Omega^3 \rightarrow \mathbb{R}$  by

$$g(\omega_1, \omega_2, \omega_3) = \begin{cases} f(\omega_1, \omega_2) & \text{if } \omega_1 \in A, \\ f(\omega_1, \omega_3) & \text{if } \omega_1 \notin A. \end{cases}$$

Then

$$\iiint_{\Omega^3} g \, d(P \times P \times P) = \iint_{\Omega^2} f \, d(P \times P).$$

Prove it.

It follows that  $Y$  is distributed like  $X$ , therefore, like  $B$ , which proves Lemma 2c3 for  $n = 0$ .

## 2e The proof, at last

If two non-overlapping changes are separately harmless, then they are jointly harmless in the following sense.

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<sup>1</sup>See also 2f4.

**2e1 Exercise.** (a) Let  $X, Y_1, Y_2$  be identically distributed random variables (on a probability space) such that  $\mathbb{P}(Y_1 \neq X \text{ and } Y_2 \neq X) = 0$ . Then the random variable  $Z$  defined by

$$Z = \begin{cases} X & \text{if } Y_1 = X \text{ and } Y_2 = X, \\ Y_1 & \text{if } Y_1 \neq X \text{ and } Y_2 = X, \\ Y_2 & \text{if } Y_1 = X \text{ and } Y_2 \neq X \end{cases}$$

is distributed like  $X$ .

(b) The same holds for random vectors and random continuous functions. Prove it.

The same holds for any finite (or countable) collection of pairwise non-overlapping changes.

*Proof of 2c3.* We consider random continuous functions

$$Y_{n,k}(t)(\omega_1, \omega_2) = \begin{cases} B(t)(\omega_1) & \text{if } t \leq k \cdot 2^{-n}, \\ B(t)(\omega_1) & \text{if } t \geq k \cdot 2^{-n} \text{ and } \tau_n(\omega_1) \neq k \cdot 2^{-n}, \\ B(k \cdot 2^{-n})(\omega_1) + B(t - k \cdot 2^{-n})(\omega_2) & \text{if } t \geq k \cdot 2^{-n} \text{ and } \tau_n(\omega_1) = k \cdot 2^{-n}. \end{cases}$$

Each  $Y_{n,k}$  is distributed like  $B$  by the argument of 2d. It remains to apply 2e1.  $\square$

## 2f Technicalities: sigma-fields and stopping times

The Borel  $\sigma$ -field<sup>1</sup>  $\mathcal{B}$  on the space  $C[0, 1]$  of all continuous functions  $[0, 1] \rightarrow \mathbb{R}$  can be defined in many equivalent ways; here is the best one for our purposes:

$\mathcal{B}$  is generated by the functions

$$(2f1) \quad C[0, 1] \ni f \mapsto f(t) \in \mathbb{R} \\ \text{where } t \text{ runs over } [0, 1].$$

**2f2 Exercise.** Prove that each of the following four sets of functions  $C[0, 1] \rightarrow \mathbb{R}$  generates the Borel  $\sigma$ -field:

- (a)  $f \mapsto f(t)$  for rational  $t \in [0, 1]$ ;
- (b)  $f \mapsto \max_{[a,b]} f(\cdot)$  for  $[a, b] \subset [0, 1]$ ;
- (c)  $f \mapsto \int_a^b f(x) dx$  for  $[a, b] \subset [0, 1]$ ;
- (d)  $f \mapsto \|f - g\|$  for  $g \in C[0, 1]$ .

Prove it.

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<sup>1</sup>In other words, “ $\sigma$ -algebra”.



It follows easily from (d) that the Borel  $\sigma$ -field is generated by open (or closed) balls, as well as by open (or closed) sets.

For any  $t \in [0, \infty)$  the Borel  $\sigma$ -field on  $C[0, t]$  is defined similarly.

Now, for a given  $t \in [0, \infty)$  we define a  $\sigma$ -field  $\mathcal{B}_t$  on the set  $C[0, \infty)$  of all continuous (not necessarily bounded) functions  $[0, \infty) \rightarrow \mathbb{R}$  as consisting of inverse images of all Borel subsets of  $C[0, t]$  under the restriction map

$$C[0, \infty) \ni f \mapsto f|_{[0, t]} \in C[0, t].$$

Clearly,  $\mathcal{B}_t$  is generated by the functions

$$C[0, \infty) \ni f \mapsto f(s) \in \mathbb{R}$$

for  $s \in [0, t]$ .

The  $\sigma$ -field generated by  $\cup_t \mathcal{B}_t$  will be denoted by  $\mathcal{B}_\infty$  and called the Borel  $\sigma$ -field of  $C[0, \infty)$ . Clearly,  $\mathcal{B}_\infty$  is generated by the functions

$$C[0, \infty) \ni f \mapsto f(t) \in \mathbb{R}$$

for  $t \in [0, \infty)$ .

Here are two equivalent definitions of a *random continuous function*.

**2f3 Exercise.** Let  $(\Omega, \mathcal{F}, P)$  be a probability space. Then the following two conditions on a function  $X : \Omega \rightarrow C[0, \infty)$  are equivalent:

(a) for each  $t \in [0, \infty)$  the function

$$\Omega \ni \omega \mapsto X(t)(\omega)$$

is  $\mathcal{F}$ -measurable;

(b) for each  $\mathcal{B}_\infty$ -measurable function  $\varphi : C[0, \infty) \rightarrow \mathbb{R}$ , the function

$$\Omega \ni \omega \mapsto \varphi(X(\cdot)(\omega))$$

is  $\mathcal{F}$ -measurable.

Prove it.

For the next exercise you need something like the monotone class theorem or Dynkin's  $\pi - \lambda$  theorem; see [1, Appendix A2, (2.1) and (2.2)].

Here are two equivalent definitions of *identically distributed* random continuous functions.

**2f4 Exercise.** The following two conditions on random continuous functions<sup>1</sup>  $X, Y$  are equivalent:

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<sup>1</sup>Maybe, on different probability spaces.

(a) for every  $n$  and  $t_1, \dots, t_n \in [0, \infty)$  the random vectors  $(X(t_1), \dots, X(t_n))$  and  $(Y(t_1), \dots, Y(t_n))$  are identically distributed;

(b) for every  $\mathcal{B}_\infty$ -measurable function  $\varphi : C[0, \infty) \rightarrow \mathbb{R}$  the random variables  $\varphi(X(\cdot))$  and  $\varphi(Y(\cdot))$  are identically distributed.

Prove it.

**2f5 Definition.** A *stopping time* is a function  $T : C[0, \infty) \rightarrow [0, \infty]$  such that

$$\{f \in C[0, \infty) : T(f) \leq t\} \in \mathcal{B}_t$$

for all  $t \in [0, \infty)$ .

**2f6 Exercise.** The hitting time  $T_1$  defined by

$$T_1(f) = \inf\{t : f(t) = 1\}$$

( $\infty$ , if the set is empty) is a stopping time.

Prove it.

**2f7 Exercise.** The function  $L$  defined by

$$L(f) = \sup\{t \in [0, 1] : f(t) = 0\}$$

(0, if the set is empty) is not a stopping time.

Prove it.

Here is the strong Markov property of the Brownian motion.

**2f8 Theorem.** If  $T$  is a stopping time then the random function

$$Y(t)(\omega_1, \omega_2) = \begin{cases} B(t)(\omega_1) & \text{for } t \leq T(\omega_1), \\ B(T(\omega_1))(\omega_1) + B(t - T(\omega_1))(\omega_2) & \text{for } t \geq T(\omega_1) \end{cases}$$

on  $\Omega \times \Omega$  is distributed like the Brownian motion  $B$ .

The proof is quite similar to the proof of 2b2.

**2f9 Remark.** A weaker (than 2f5) assumption

$$\{f \in C[0, \infty) : T(f) < t\} \in \mathcal{B}_t \quad \text{for all } t \in [0, \infty)$$

is still sufficient for Theorem 2f8 to hold.

(Anyway, a delay is stipulated by the proof, recall 2c). Such  $T$  is called a stopping time of the (right-continuous) filtration  $(\mathcal{B}_{t+})_t$ , where  $\mathcal{B}_{t+} = \bigcap_{\varepsilon > 0} \mathcal{B}_{t+\varepsilon}$ . In contrast, 2f5 defines a stopping time of the filtration  $(\mathcal{B}_t)_t$ . (Generally, a filtration is defined as an increasing family of  $\sigma$ -fields.)

Here is an example of a stopping time of  $(\mathcal{B}_{t+})_t$  but not  $(\mathcal{B}_t)_t$ :

$$T_{1+} = \inf\{t : B(t) > 1\}.$$

Note that  $T_t \downarrow T_{1+}$  as  $t \downarrow 1$ . Similarly,  $T_{x+}$  are introduced for all  $x \in [0, \infty)$ . Due to 2f9, all said in 2b about the process  $(T_x)_{x \in [0, \infty)}$  holds also for  $(T_{x+})_{x \in [0, \infty)}$ , except for the left continuity; this time we get right continuity.

The random function  $(T_{x+})_{x \in [0, \infty)}$  is distributed as the special Lévy process.

**2f10 Exercise.**  $\mathbb{P}(T_x = T_{x+}) = 1$  for each  $x \in [0, \infty)$ .

Prove it.

## 2g Hints to exercises

2a2: Calculate the joint distribution of  $Y(t_1), Y(t_2) - Y(t_1), \dots, Y(t_n) - Y(t_{n-1})$  assuming that  $t_1 < \dots < t_n$  and  $1 \in \{t_1, \dots, t_n\}$ .

2a4:  $\{\omega : T_x > t\} = \{\omega : \sup_{[0,t]} B(\cdot) < x\}$ .

2a5: Use 1c2.

2b3:  $\mathbb{P}(\max_{[0,t]} B(\cdot) \geq 1) = \mathbb{P}(T_1 \leq t)$ ; use 2b2.

2b5:  $\mathbb{P}(T_x < \infty) = 1$ .

2b6:  $\lim_{x \rightarrow 0+} \mathbb{P}(T_x < \varepsilon) = ?$

2b7: use 2b6.

2b10: Use 2b2.

2d1: Fubini theorem.

2f2: (a), (b), (c): if  $\varphi_n$  are measurable (w.r.t. a given  $\sigma$ -field) and  $\varphi_n \rightarrow \varphi$  pointwise, then  $\varphi$  is also measurable. (d): take a sequence  $(g_n)_n$  dense in  $C[0, 1]$  and note that  $\sup_{n: \|f - g_n\| < 1} g_n(t) = f(t) + 1$ .

2f3: (a)  $\implies$  (b): all sets  $A \subset C[0, \infty)$  such that  $X^{-1}(A) \in \mathcal{F}$  are a  $\sigma$ -field.

2f10:  $T_x$  and  $T_{x+}$  are identically distributed, and  $T_x \leq T_{x+}$ .

## References

- [1] R. Durrett, *Probability: theory and examples*, 1996.

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